Markov Chain Monte Carlo (MCMC) Simulation

Outline

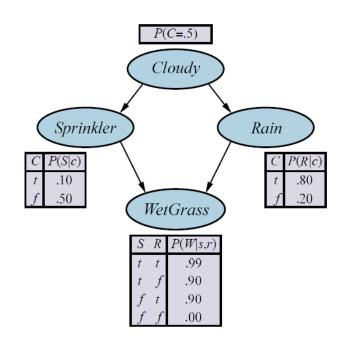
- I. Gibbs simulation
- II. Posterior distribution under the Markov blanket
- III. Markov chains

^{*} Figures are from the <u>textbook site</u>.

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- I. Gibbs simulation
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Use of Markov Chains

- A Markov Chain Monte Carlo (MCMC) algorithm
 - specifies a value for every variable at the current state (i.e., sample), and
 - generates a next state by making random changes to the current state.

 Markov chain is a random process that generates a sequence of states.

An MCMC algorithm is well suited for Bayes nets that

- starts with an arbitrary state
- fix evidence variables at their observed values
- chose a variable X_i out of the m nonevidence variables with a specified probability:

$$\rho(i) = P(X_i \text{ is chosen among } X_1, \dots, X_m)$$

randomly sample a value for the chosen variable X_i

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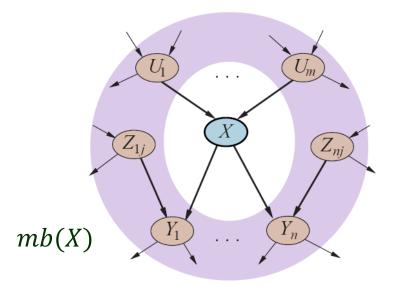
 X_i is independent of all the variables outside of its *Markov blanket* $mb(X_i)$.

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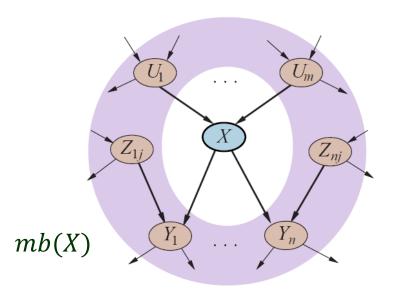
 X_i 's parents, children, and children's other parents

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- chose a variable X_i out of the m nonevidence variables with a specified probability:

$$\rho(i) = P(X_i \text{ is chosen among } X_1, \dots, X_m)$$

• randomly sample a value for the chosen variable X_i according to



```
P(X_i \mid mb(X_i)) // how to compute? // (described later)
```

 X_i is independent of all the variables outside of its *Markov blanket* $mb(X_i)$.

 X_i 's parents, children, and children's other parents

The Algorithm

```
function GIBBS-ASK(X, \mathbf{e}, bn, N) returns an estimate of \mathbf{P}(X \mid \mathbf{e}) local variables: \mathbf{C}, a vector of counts for each value of X, initially zero \mathbf{Z}, the nonevidence variables in bn //X \in \mathbf{Z} \mathbf{x}, the current state of the network, initialized from \mathbf{e} N, number of samples initialize \mathbf{x} with random values for the variables in \mathbf{Z} for k = 1 to N do choose any variable Z_i from \mathbf{Z} according to any distribution \rho(i) set the value of Z_i in \mathbf{x} by sampling from \mathbf{P}(Z_i \mid mb(Z_i)) \mathbf{C}[j] \leftarrow \mathbf{C}[j] + 1 where x_j is the value of X in \mathbf{x} return NORMALIZE(\mathbf{C})
```

The Algorithm

function GIBBS-ASK(X, \mathbf{e} , bn, N) returns an estimate of $\mathbf{P}(X \mid \mathbf{e})$ local variables: \mathbf{C} , a vector of counts for each value of X, initially zero \mathbf{Z} , the nonevidence variables in bn $//X \in \mathbf{Z}$ \mathbf{x} , the current state of the network, initialized from \mathbf{e} N, number of samples initialize \mathbf{x} with random values for the variables in \mathbf{Z} for k = 1 to N do choose any variable Z_i from \mathbf{Z} according to any distribution $\rho(i)$ set the value of Z_i in \mathbf{x} by sampling from $\mathbf{P}(Z_i \mid mb(Z_i))$ $\mathbf{C}[j] \leftarrow \mathbf{C}[j] + 1$ where x_j is the value of X in \mathbf{x} return NORMALIZE(\mathbf{C})

Values of
$$X$$
: $\begin{bmatrix} x_1 & \cdots & x_j & \cdots & x_m \\ & & & & \\ C & & & & & \\ 1 & \cdots & j & \cdots & m \end{bmatrix}$

The Algorithm

function GIBBS-ASK(X, \mathbf{e} , bn, N) **returns** an estimate of $\mathbf{P}(X \mid \mathbf{e})$

local variables: C, a vector of counts for each value of X, initially zero

Z, the nonevidence variables in $bn // X \in \mathbf{Z}$

x, the current state of the network, initialized from e

N, number of samples

initialize x with random values for the variables in Z

for k = 1 to N do

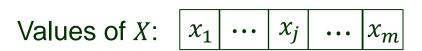
choose any variable Z_i from **Z** according to any distribution $\rho(i)$ set the value of Z_i in **x** by sampling from $\mathbf{P}(Z_i \mid mb(Z_i))$

 $\mathbf{C}[j] \leftarrow \mathbf{C}[j] + 1$ where x_j is the value of X in \mathbf{x}

return NORMALIZE(C)

Two cases in iteration k:

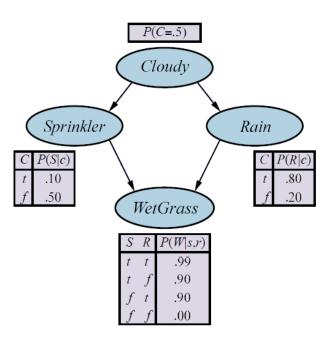
- $Z_i = X$. x_j is the newly sampled value of X. The index j changes from the previous iteration only if the value x_i does.
- Z_i ≠ X. The value of X does not change, neither does j. The same counter C[j] as in the previous iteration is incremented again.



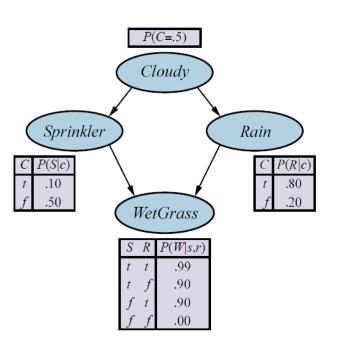


Gibbs sampling for X_i is conditioned on the current values of the variables in its Markov blanket.

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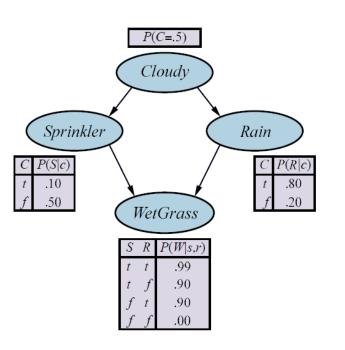


Gibbs sampling for X_i is conditioned on the current values of the variables in its Markov blanket.



Query $P(Rain \mid Sprinkler = true, WetGrass = true)$

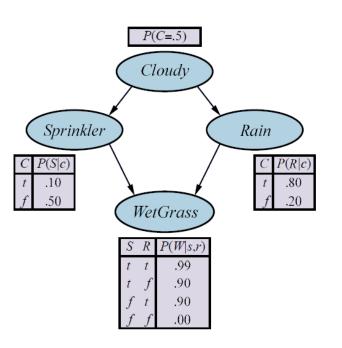
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Query $P(Rain \mid Sprinkler = true, WetGrass = true)$

Initial state [true, true, false, true]

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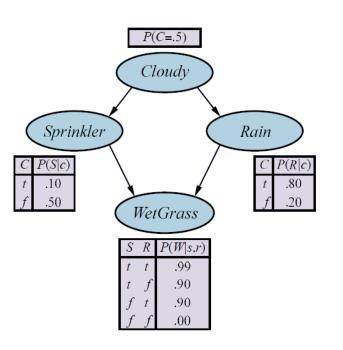


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Initial state [true,true,false,true]

evidence variables *Sprinkler* and *WetGrass* fixed to their observed values

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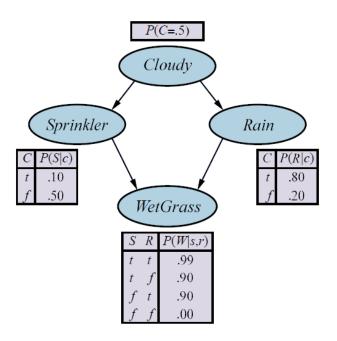
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randomly generated values for nonevidence variables *Cloudy* and *Rain*

Initial state [true, true, false, true]

evidence variables *Sprinkler* and *WetGrass* fixed to their observed values

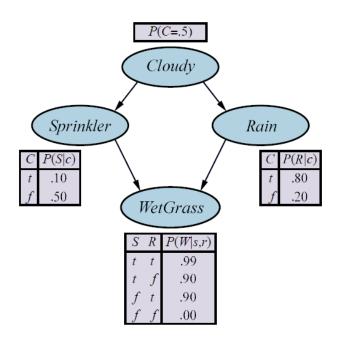
Query $P(Rain \mid Sprinkler = true, WetGrass = true)$



Order: Cloudy, Sprinkler, Rain, WetGrass

[true,true,false,true] (initial state)

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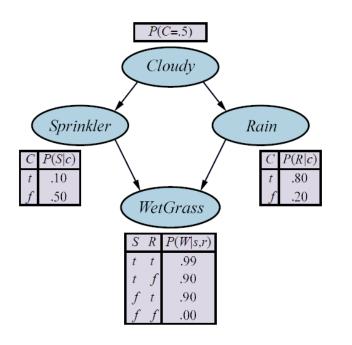


• Non-evidence variables are then sampled in random order following some probability distribution $\rho(i)$.

Order: Cloudy, Sprinkler, Rain, WetGrass

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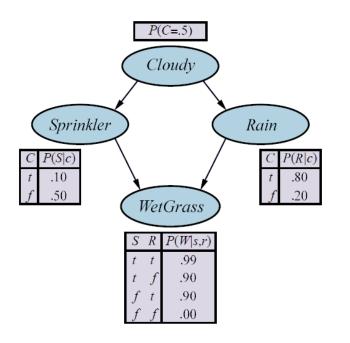


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 - ♠ Cloudy is chosen and sampled given the current values of its Markov blanket {Sprinkler, Rain}.

Order: Cloudy, Sprinkler, Rain, WetGrass

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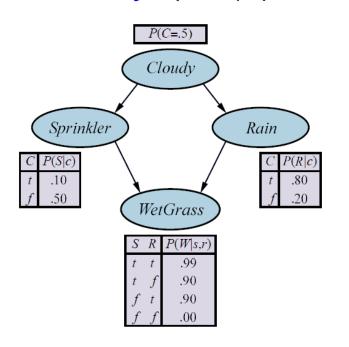
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 - Sampling distribution:

P(Cloudy | Sprinkler = true, Rain = false)

Order: Cloudy, Sprinkler, Rain, WetGrass

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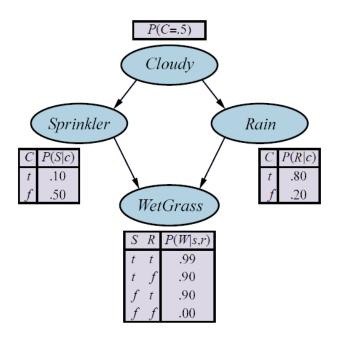


How to calculate Markov blanket distribution $P(X_i | mb(X_i))$?

Order: Cloudy, Sprinkler, Rain, WetGrass

[true,true,false,true] · (initial state)

Query $P(Rain \mid Sprinkler = true, WetGrass = true)$



Order: Cloudy, Sprinkler, Rain, WetGrass

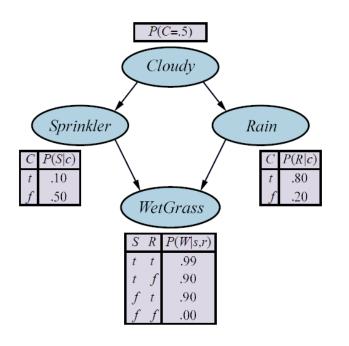
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Sampling result: Cloudy = false.

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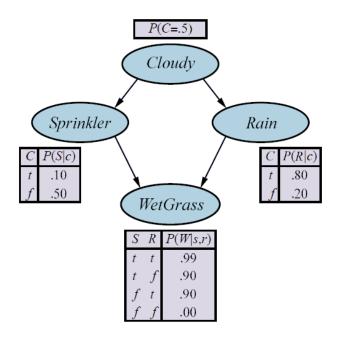
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$$[true, true, false, true] \rightarrow [false, true, false, true]$$

(initial state) (2nd state)

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Query $P(Rain \mid Sprinkler = true, WetGrass = true)$



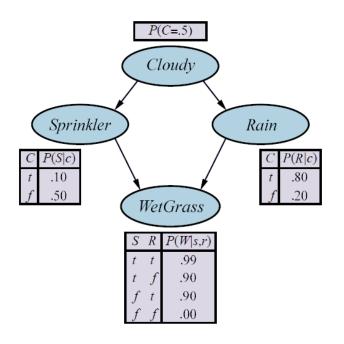
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 How to calculate Markov blanket distribution P(Xi | mb(Xi))?

 $[true, true, false, true] \rightarrow [false, true, false, true]$ (initial state) (2nd state)

Query $P(Rain \mid Sprinkler = true, WetGrass = true)$



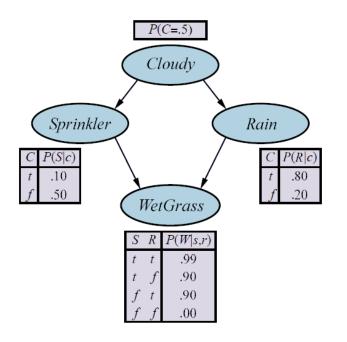
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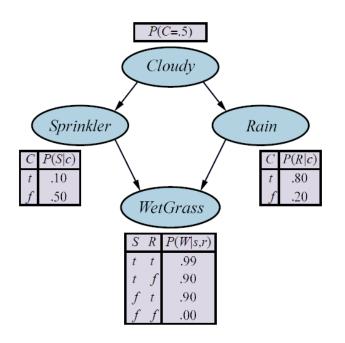
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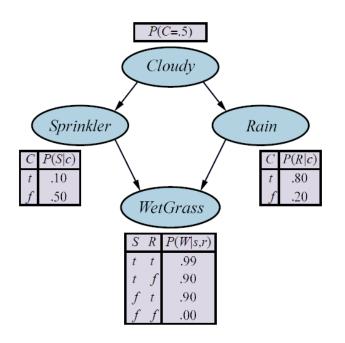
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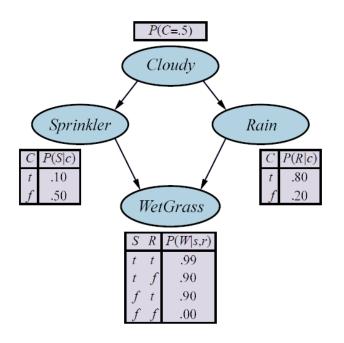
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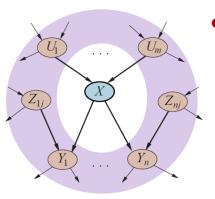
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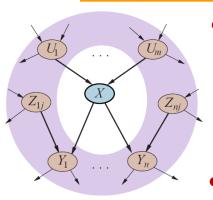
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- $[true, true, false, true] \rightarrow [false, true, false, true]$ $(initial state) \qquad (2^{nd} state)$ $\rightarrow [false, true, true, true]$ $(3^{rd} state)$ $\rightarrow \cdots$

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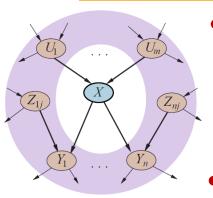
- Sampling result: Cloudy = false.
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 How to calculate Markov blanket
 - Sampling distribution: distribution $P(X_i \mid mb(X_i))$?



- $MB(X_i)$: variables in the Markov blanket of X_i .
 - Parents(X_i): parents of X_i (e.g., U_1 , ..., U_m)
 - Children (X_i) : children of X_i (e.g., $Y_1, ..., Y_n$)
 - Others(X_i): other parents of X_i 's children (e.g, Z_{1j} , ..., Z_{nj} , ...)



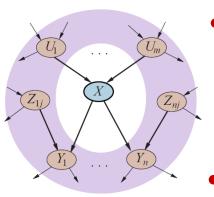
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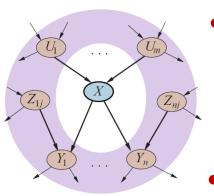
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We want to compute the probability distribution $P(X_i \mid mb(X_i))$.

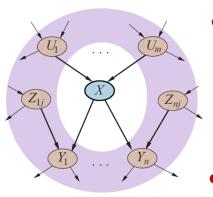
 $P(X_i \mid mb(X_i)) = P(X_i \mid parents(X_i), children(X_i), others(X_i))$



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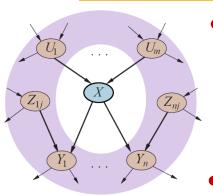


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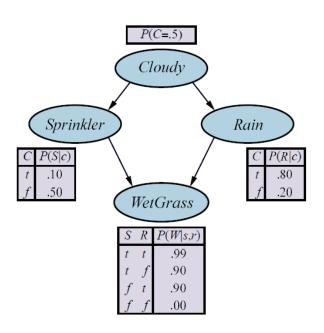
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\begin{aligned} \textbf{\textit{P}}\big(X_i \mid mb(X_i)\big) &= \textbf{\textit{P}}(X_i \mid parents(X_i), children(X_i), others(X_i)) \\ &= \alpha \textbf{\textit{P}}\left(X_i, parents(X_i), children(X_i), others(X_i)\right) \\ &= \alpha \textbf{\textit{P}}(X_i \mid parents(X_i), others(X_i)) \cdot \textbf{\textit{P}}\big(children(X_i) \mid parents(X_i), X_i, others(X_i)\big) \\ &= \alpha \textbf{\textit{P}}(X_i \mid parents(X_i)) \cdot \textbf{\textit{P}}\big(children(X_i) \mid X_i, others(X_i)\big) \end{aligned}
```



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$$P(X_i \mid mb(X_i)) = \alpha P(X_i \mid parents(X_i)) \qquad \prod_{Y_j \in Children(X_i)} P(y_j \mid parents(Y_j))$$

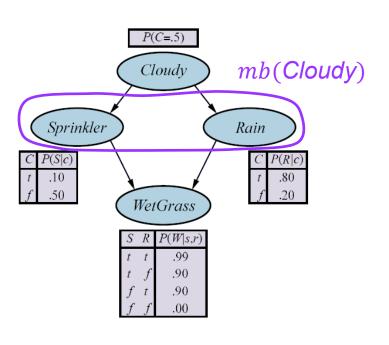
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Sampling distribution:

P(Cloudy | Sprinkler = true, Rain = false)

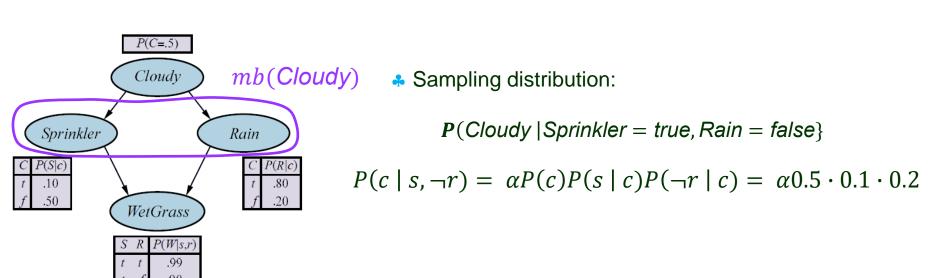
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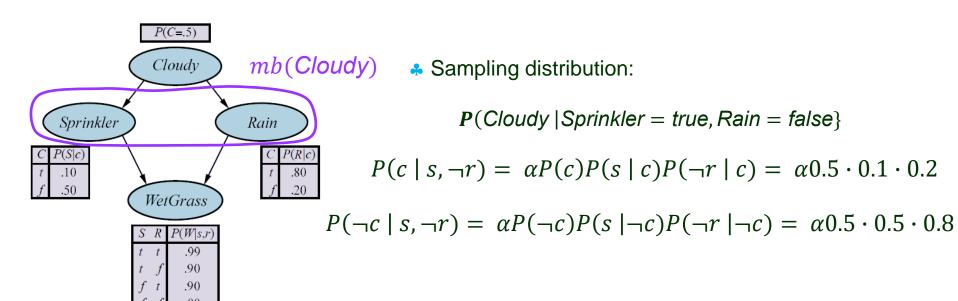
Sampling distribution:

P(Cloudy | Sprinkler = true, Rain = false)

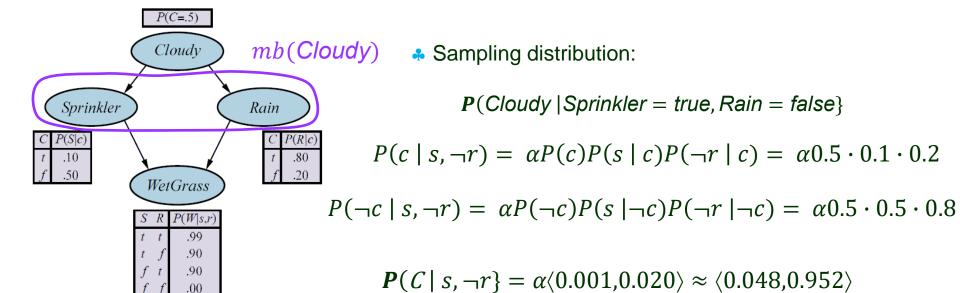
$$P(X_i \mid mb(X_i)) = \alpha P(X_i \mid parents(X_i)) \prod_{Y_j \in Children(X_i)} P(y_j \mid parents(Y_j))$$



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Query $P(Rain \mid Sprinkler = true, WetGrass = true)$

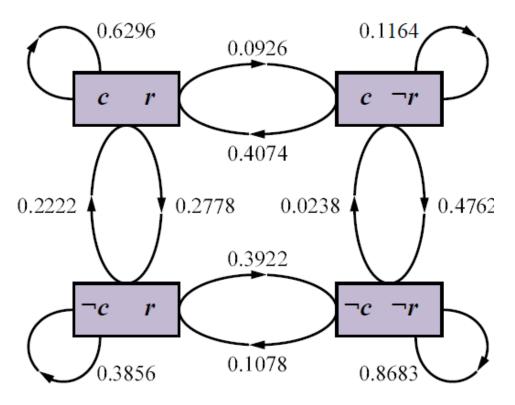
A state *need only include all nonevidence variables*, for example, $\{c, r\}$.

 $mb(Rain) = \{Cloudy, Sprinkler, WetGrass\}$

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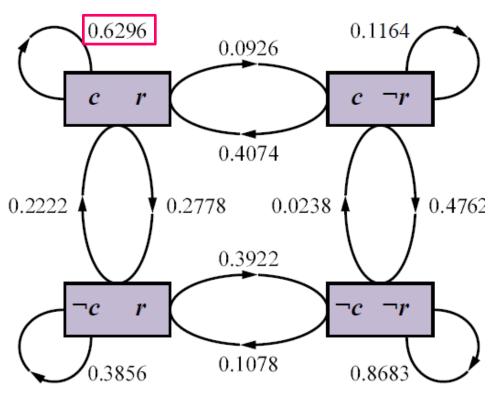
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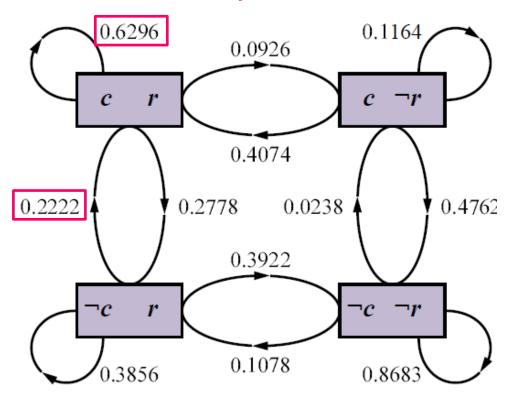


$$0.6296 = \rho(Cloudy) \cdot P(c \mid r, s) + \rho(Rain) \cdot P(r \mid c, s, w)$$

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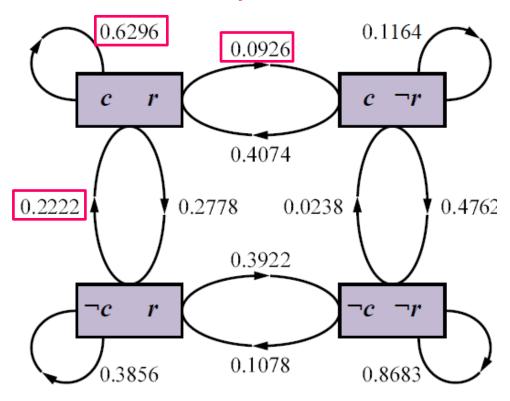


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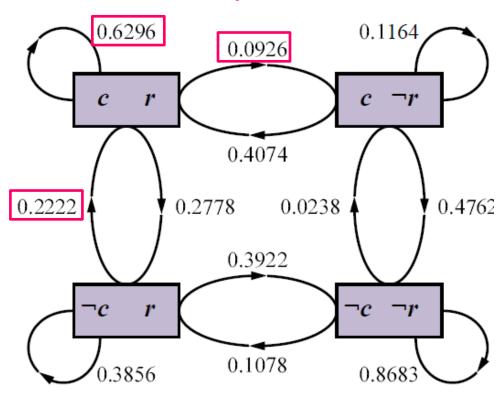


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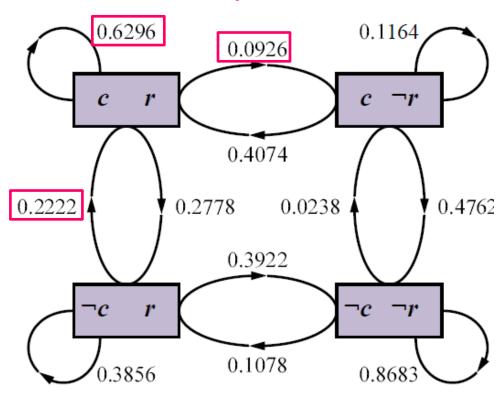
Markov chain from uniform choice of the two nonevidence variables ($\rho(Cloudy) = \rho(Rain) = 0.5$)

Probabilities with all the outgoing links of each node sum to 1, e.g., 0.6296 + 0.0926 + 0.2778 = 1.

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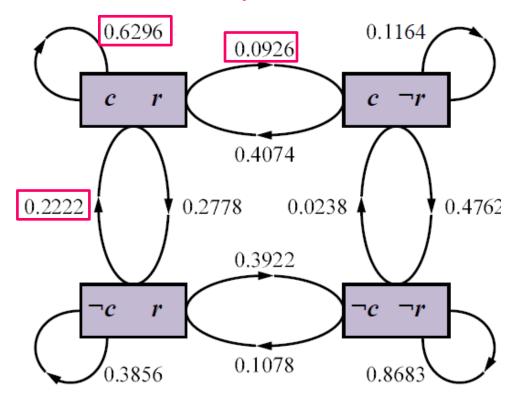
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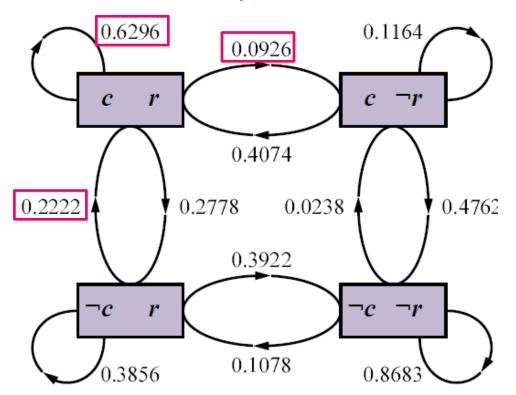
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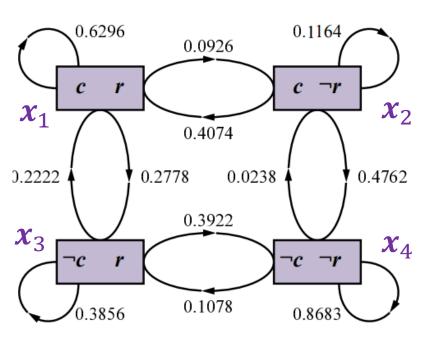
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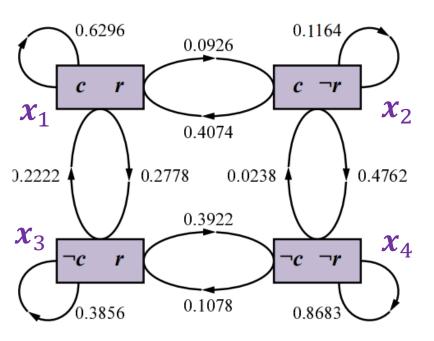
If the process visits 20 states with Rain = true and 60 states with Rain = false, then the answer to the query is $\alpha \langle 20,60 \rangle = \langle 0.25,0.75 \rangle$.

Why does Gibbs sampling work? Or, why does its estimates converge to correct values in the limit?



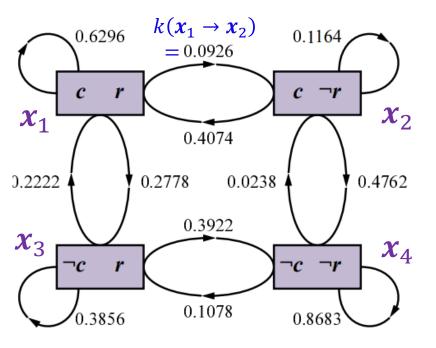
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Transition kernel k assigns a probability $k(x \rightarrow x')$ to every transition from a state x to another state x' in the Markov chain.



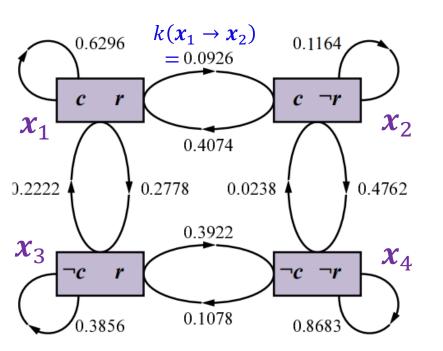
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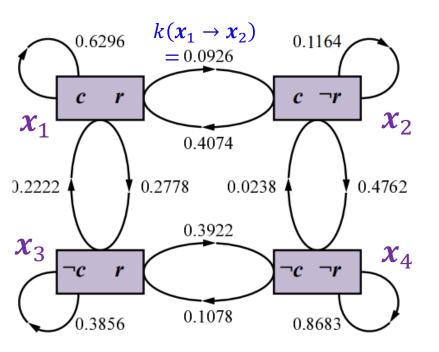
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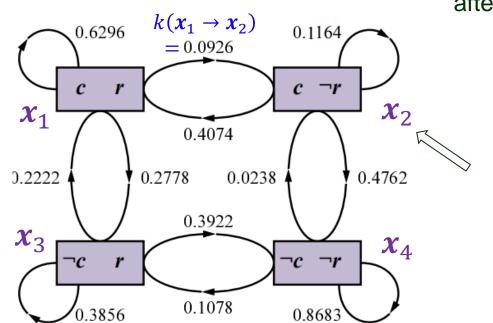


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$$\pi_{t+1}(x') = \sum_{x} \pi_t(x) k(x \to x')$$

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$$\pi_{t+1}(\mathbf{x}') = \sum_{\mathbf{x}} \pi_t(\mathbf{x}) k(\mathbf{x} \to \mathbf{x}')$$

$$\pi_{t+1}(\boldsymbol{x}_2) = 0.0926 \, \pi_t(\boldsymbol{x}_1) + 0.1164 \pi_t(\boldsymbol{x}_2) \\ + 0.0238 \, \pi_t(\boldsymbol{x}_4)$$

The chain has reached its *stationary* distribution if $\pi_{t+1}(x) = \pi_t(x)$ for all x. We then call this stationary distribution π .

X

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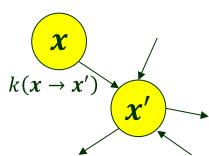
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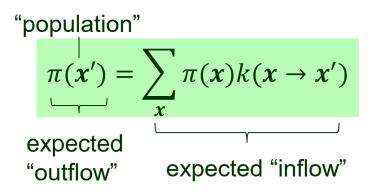
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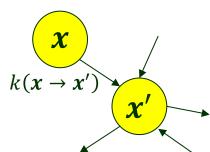
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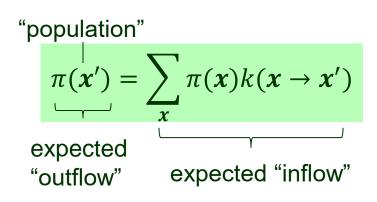
- a) every state is reachable from every other state, and
- b) there exists *exactly one* stationary distribution π .

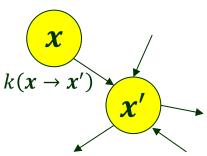
In a stationary distribution π , the expected "outflow" from each state is equal to the expected "inflow" from all the other states.





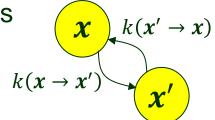
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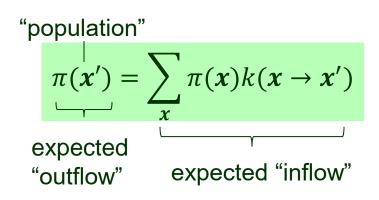


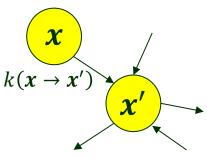
A *detailed balance* k with π is a distribution that satisfies

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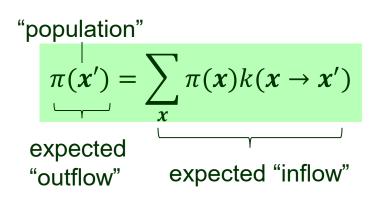


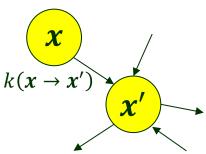
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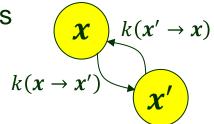
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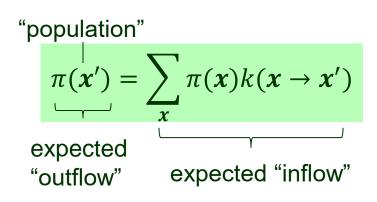
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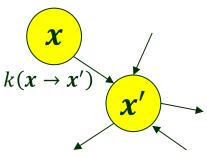
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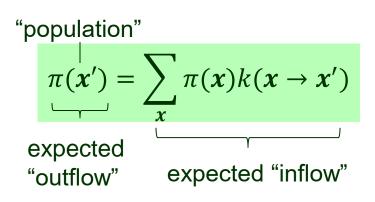
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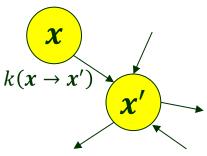
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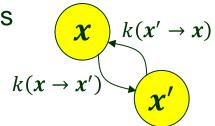
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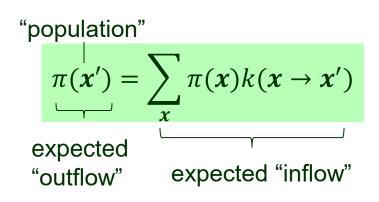
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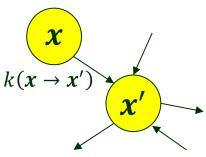
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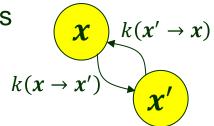
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